The ilc package: Iterative Lee-Carter

Han Lin Shang

Research School of Finance, <u>Actuarial Studies</u> and Applied Statistics, Australian National University

> Presentation prepared for the useR2015 Aalborg, 1st July 2015

Acknowledgement: Zoltan Butt and Steven Haberman (Cass Business School, London)

Objectives of methods and ilc package Model, estimation and forecasting Demonstration Conclusion

Motivating example: UK male log mortality rates

Force of mortality is

$$\widehat{\mu}_{x,t} = \widehat{m}_{x,t} = \frac{y_{x,t}}{e_{x,t}}$$

where $y_{x,t}$ and $e_{x,t}$ represent the number of deaths and corresponding central exposure for any given age group at year t. Obtain a $n \times p$ matrix to represent age and time dimensions

LC model is defined as

$$\ln(m_{x,t}) = \alpha_x + \beta_x \kappa_t + \varepsilon_{x,t}$$

1 α_x represents a constant age-specific pattern

Objectives of methods and ilc package Model, estimation and forecasting Demonstration Conclusion

LC model is defined as

$$\ln(m_{x,t}) = \alpha_x + \beta_x \kappa_t + \varepsilon_{x,t}$$

1 α_x represents a constant age-specific pattern **2** κ_t measures the trend in mortality over time

LC model is defined as

$$\ln(m_{x,t}) = \alpha_x + \beta_x \kappa_t + \varepsilon_{x,t}$$

- **1** α_x represents a constant age-specific pattern
- **2** κ_t measures the trend in mortality over time
- 3 β_x measures the age-specific deviations of mortality change from the overall trend

LC model is defined as

$$\ln(m_{x,t}) = \alpha_x + \beta_x \kappa_t + \varepsilon_{x,t}$$

- **1** α_x represents a constant age-specific pattern
- **2** κ_t measures the trend in mortality over time
- 3 β_x measures the age-specific deviations of mortality change from the overall trend
- 4 $\varepsilon_{x,t}$ are assumed to be $N\left(0,\sigma^2\right)$ random effects by age and time

 Implement an iterative regression method for analysing age-period mortality rates via generalised Lee-Carter (LC) model

- Implement an iterative regression method for analysing age-period mortality rates via generalised Lee-Carter (LC) model
- Use Generalised Linear Model (GLM) model (Renshaw and Haberman 2006)

- Implement an iterative regression method for analysing age-period mortality rates via generalised Lee-Carter (LC) model
- Use Generalised Linear Model (GLM) model (Renshaw and Haberman 2006)
- Develop and implement a stratified LC model for the measurement of the additive effect on the log scale of an explanatory factor (other than age and time)

- Implement an iterative regression method for analysing age-period mortality rates via generalised Lee-Carter (LC) model
- Use Generalised Linear Model (GLM) model (Renshaw and Haberman 2006)
- Develop and implement a stratified LC model for the measurement of the additive effect on the log scale of an explanatory factor (other than age and time)
- Produce forecasts of age-specific mortality rates and life expectancy

Extend LC model based on the Gaussian error structure to Poisson

Objectives of methods and ilc package Model, estimation and forecasting Demonstration Conclusion

- Extend LC model based on the <u>Gaussian</u> error structure to <u>Poisson</u>
- Instead of singular value decomposition, consider a regression model based on <u>Poisson likelihood maximisation</u>

- Extend LC model based on the <u>Gaussian</u> error structure to <u>Poisson</u>
- Instead of singular value decomposition, consider a regression model based on <u>Poisson likelihood maximisation</u>
- ilc package contains methods for the analysis of a class of six log-linear models (capturing age, period, cohort) in the GLM with Poisson errors

To assess goodness of fit of the regression, estimation routines support a range of residual diagnostic plots

- To assess goodness of fit of the regression, estimation routines support a range of residual diagnostic plots
- Allows preliminary data corrections, to replace missing data cells, but also to eliminate potential outliers that might result from data inaccuracies

- To assess goodness of fit of the regression, estimation routines support a range of residual diagnostic plots
- Allows preliminary data corrections, to replace missing data cells, but also to eliminate potential outliers that might result from data inaccuracies
- 3 Includes two simple methods of "closing-out" produces to correct the original data at very old ages before the application of the model

- To assess goodness of fit of the regression, estimation routines support a range of residual diagnostic plots
- Allows preliminary data corrections, to replace missing data cells, but also to eliminate potential outliers that might result from data inaccuracies
- Includes two simple methods of "closing-out" produces to correct the original data at very old ages before the application of the model
- 4 ilc package integrates with the *demography* and *forecast* packages

- To assess goodness of fit of the regression, estimation routines support a range of residual diagnostic plots
- Allows preliminary data corrections, to replace missing data cells, but also to eliminate potential outliers that might result from data inaccuracies
- Includes two simple methods of "closing-out" produces to correct the original data at very old ages before the application of the model
- 4 ilc package integrates with the *demography* and *forecast* packages
- ilc package has improved inspection and graphical visualisation of mortality data and regression output

Lee-Carter model under Poisson error

LC parameters can be estimated by maximum likelihood methods based on Poisson error distribution

Lee-Carter model under Poisson error

- LC parameters can be estimated by maximum likelihood methods based on Poisson error distribution
- 2 Assuming that age- and period-specific number of deaths are independent realisations from a Poisson distribution with parameters

$$\mathsf{E}[Y_{x,t}] = e_{x,t}\mu_{x,t}, \quad \mathsf{Var}[Y_{x,t}] = \phi\mathsf{E}[Y_{x,t}]$$

where ϕ is a measure of over-dispersion to allow for heterogeneity

Lee-Carter model under Poisson error

- LC parameters can be estimated by maximum likelihood methods based on Poisson error distribution
- 2 Assuming that age- and period-specific number of deaths are independent realisations from a Poisson distribution with parameters

$$\mathsf{E}[Y_{x,t}] = e_{x,t}\mu_{x,t}, \quad \mathsf{Var}[Y_{x,t}] = \phi\mathsf{E}[Y_{x,t}]$$

where ϕ is a measure of over-dispersion to allow for heterogeneity

3 GLM model of the response variable $Y_{x,t}$ with log-link and non-linear parameterized predictor:

$$\eta_{x,t} = \ln(\hat{y}_{x,t}) = \underbrace{\ln(e_{x,t})}_{\text{offset}} + \alpha_x + \beta_x \kappa_t$$

Maximum likelihood point estimates under the GLM approach are obtained at the minimum value of the total deviation, given by

$$D\left(y_{x,t}, \widehat{y}_{x,t}\right) = \sum_{x,t} \operatorname{dev}(x,t) = \sum_{x,t} 2\omega_{x,t} \left\{ y_{x,t} \ln \frac{y_{x,t}}{\widehat{y}_{x,t}} - (y_{x,t} - \widehat{y}_{x,t}) \right\}$$
(1)

Maximum likelihood point estimates under the GLM approach are obtained at the minimum value of the total deviation, given by

$$D\left(y_{x,t},\widehat{y}_{x,t}\right) = \sum_{x,t} \operatorname{dev}(x,t) = \sum_{x,t} 2\omega_{x,t} \left\{ y_{x,t} \ln \frac{y_{x,t}}{\widehat{y}_{x,t}} - (y_{x,t} - \widehat{y}_{x,t}) \right\}$$
(1)

where ${\rm dev}(x,t)$ are the deviance residuals that depend on a set of prior weights $\omega_{x,t}$

2 Resort to an iterative Newton-Raphson method applied to the deviance function (1) . We use the iterative procedure:

Maximum likelihood point estimates under the GLM approach are obtained at the minimum value of the total deviation, given by

$$D(y_{x,t}, \widehat{y}_{x,t}) = \sum_{x,t} \operatorname{dev}(x,t) = \sum_{x,t} 2\omega_{x,t} \left\{ y_{x,t} \ln \frac{y_{x,t}}{\widehat{y}_{x,t}} - (y_{x,t} - \widehat{y}_{x,t}) \right\}$$
(1)

- **2** Resort to an iterative Newton-Raphson method applied to the deviance function (1) . We use the iterative procedure:
 - **1** Set starting values $\widehat{\beta}_x$

Maximum likelihood point estimates under the GLM approach are obtained at the minimum value of the total deviation, given by

$$D(y_{x,t}, \hat{y}_{x,t}) = \sum_{x,t} \operatorname{dev}(x,t) = \sum_{x,t} 2\omega_{x,t} \left\{ y_{x,t} \ln \frac{y_{x,t}}{\hat{y}_{x,t}} - (y_{x,t} - \hat{y}_{x,t}) \right\}$$
(1)

- **2** Resort to an iterative Newton-Raphson method applied to the deviance function (1) . We use the iterative procedure:
 - **1** Set starting values $\hat{\beta}_x$
 - **2** Given $\hat{\beta}_x$, update $\hat{\alpha}_x$ and $\hat{\kappa}_t$

Maximum likelihood point estimates under the GLM approach are obtained at the minimum value of the total deviation, given by

$$D(y_{x,t}, \hat{y}_{x,t}) = \sum_{x,t} \operatorname{dev}(x,t) = \sum_{x,t} 2\omega_{x,t} \left\{ y_{x,t} \ln \frac{y_{x,t}}{\hat{y}_{x,t}} - (y_{x,t} - \hat{y}_{x,t}) \right\}$$
(1)

- Resort to an iterative Newton-Raphson method applied to the deviance function (1). We use the iterative procedure:
 - **1** Set starting values $\hat{\beta}_x$
 - 2 Given β_x , update $\widehat{\alpha}_x$ and $\widehat{\kappa}_t$
 - **3** Given $\widehat{\kappa}_t$, update $\widehat{\alpha}_x$ and β_x

Maximum likelihood point estimates under the GLM approach are obtained at the minimum value of the total deviation, given by

$$D(y_{x,t}, \hat{y}_{x,t}) = \sum_{x,t} \operatorname{dev}(x,t) = \sum_{x,t} 2\omega_{x,t} \left\{ y_{x,t} \ln \frac{y_{x,t}}{\hat{y}_{x,t}} - (y_{x,t} - \hat{y}_{x,t}) \right\}$$
(1)

- **2** Resort to an iterative Newton-Raphson method applied to the deviance function (1) . We use the iterative procedure:
 - **1** Set starting values $\hat{\beta}_x$
 - 2 Given $\widehat{\beta}_x$, update $\widehat{\alpha}_x$ and $\widehat{\widehat{\kappa}}_t$
 - 3 Given $\widehat{\kappa}_t$, update $\widehat{\alpha}_x$ and $\widehat{\beta}_x$
 - **4** Compute $D(y_{x,t}, \hat{y}_{x,t})$

Maximum likelihood point estimates under the GLM approach are obtained at the minimum value of the total deviation, given by

$$D\left(y_{x,t},\widehat{y}_{x,t}\right) = \sum_{x,t} \operatorname{dev}(x,t) = \sum_{x,t} 2\omega_{x,t} \left\{ y_{x,t} \ln \frac{y_{x,t}}{\widehat{y}_{x,t}} - (y_{x,t} - \widehat{y}_{x,t}) \right\}$$
(1)

- **2** Resort to an iterative Newton-Raphson method applied to the deviance function (1) . We use the iterative procedure:
 - **1** Set starting values $\widehat{\beta}_x$
 - **2** Given $\widehat{\beta}_x$, update $\widehat{\alpha}_x$ and $\widehat{\kappa}_t$
 - 3 Given $\widehat{\kappa}_t$, update $\widehat{\alpha}_x$ and $\widehat{\beta}_x$
 - **4** Compute $D(y_{x,t}, \widehat{y}_{x,t})$
 - **5** Repeat the updating cycle; stop when $D(y_{x,t}, \hat{y}_{x,t})$ converges

Basic LC model can be extended to include an additional bilinear term, containing a second period effect or a cohort effect

- Basic LC model can be extended to include an additional bilinear term, containing a second period effect or a cohort effect
- 2 Force of mortality by a generalised structure is given as

$$\mu_{x,t} = \exp\left(\alpha_x + \underline{\beta_x^{(0)}}l_{t-x} + \beta_x^{(1)}\kappa_t\right)$$

where α_x : main age profile; l_{t-x} : cohort effect; κ_t : period

Lee-Carter model with additional covariates

Additional factor depends on the size and nature of the mortality experience, such as geographical, <u>socio-economic</u> or <u>race differences</u>

Lee-Carter model with additional covariates

- Additional factor depends on the size and nature of the mortality experience, such as geographical, <u>socio-economic</u> or <u>race differences</u>
- 2 Consider a cross-classified mortality experience observed over age x, period t and an extra variate g made up of $(k \times n \times l)$ data cells

Lee-Carter model with additional covariates

- Additional factor depends on the size and nature of the mortality experience, such as geographical, <u>socio-economic</u> or <u>race differences</u>
- 2 Consider a cross-classified mortality experience observed over age x, period t and an extra variate g made up of $(k \times n \times l)$ data cells
- 3 Stratified LC model is given by

$$\eta_{x,t,g} = \ln(\widehat{y}_{x,t,g}) = \underbrace{\ln(e_{x,t,g})}_{\text{offset}} + \alpha_x + \underline{\alpha_g} + \beta_x \kappa_t,$$

where α_g measures the relative differences between the age-specific log mortality profiles among subgroups defined by the extra variate g

Forecasting

1 Forecasting mortality in the LC family of models is based on time series prediction of the calendar time dependent parameters (l_{t-x}, κ_t)

Forecasting

- **I** Forecasting mortality in the LC family of models is based on time series prediction of the calendar time dependent parameters (l_{t-x}, κ_t)
- 2 Mortality rate forecasts can be written as

$$\widehat{\mu}_{x,n+h} = \exp\left(\widehat{\alpha}_x + \widehat{\beta}_x^{(0)}\widehat{l}_{n+h-x} + \widehat{\beta}_x^{(1)}\widehat{\kappa}_{n+h}\right)$$

where \widehat{l}_{n+h-x} and $\widehat{\kappa}_{n+h}$ represent the forecast cohort and period effects

Forecasting

- Forecasting mortality in the LC family of models is based on time series prediction of the calendar time dependent parameters (l_{t-x}, κ_t)
- 2 Mortality rate forecasts can be written as

$$\widehat{\mu}_{x,n+h} = \exp\left(\widehat{\alpha}_x + \widehat{\beta}_x^{(0)}\widehat{l}_{n+h-x} + \widehat{\beta}_x^{(1)}\widehat{\kappa}_{n+h}\right)$$

where \widehat{l}_{n+h-x} and $\widehat{\kappa}_{n+h}$ represent the forecast cohort and period effects

3 Random walk with drift, ARIMA(0,1,0), is used to forecast period effect (κ_t), expressed as

$$\kappa_t = \kappa_{t-1} + d + e_t$$

where d measures the drift and e_t represents the white noise

Forecasting

- Forecasting mortality in the LC family of models is based on time series prediction of the calendar time dependent parameters (l_{t-x}, κ_t)
 Martality muta forecasts can be a fitted as
- 2 Mortality rate forecasts can be written as

$$\widehat{\mu}_{x,n+h} = \exp\left(\widehat{\alpha}_x + \widehat{\beta}_x^{(0)}\widehat{l}_{n+h-x} + \widehat{\beta}_x^{(1)}\widehat{\kappa}_{n+h}\right)$$

where \widehat{l}_{n+h-x} and $\widehat{\kappa}_{n+h}$ represent the forecast cohort and period effects

3 Random walk with drift, ARIMA(0,1,0), is used to forecast period effect (κ_t), expressed as

$$\kappa_t = \kappa_{t-1} + d + e_t$$

where d measures the drift and e_t represents the white noise

4 In the cohort effects, forecasts revert to the fitted parameters when h falls within the available data range

CMI data contains the mortality experience of male life office pensioners retiring at or after normal retirement age

- CMI data contains the mortality experience of male life office pensioners retiring at or after normal retirement age
- 2 Data is made up of observed central exposure and deaths for ages 50-108 from 1983 to 2003

R demo of explanatory plots

- Plot mortality rates, population-at-risk, death counts for any age group and year
 - >insp.dd(dd.cmi.pens,age=50:80,year=1985:1990)
 >insp.dd(dd.cmi.pens,what='pop',age=70:100,year=1988:1993)
 >insp.dd(dd.cmi.pens,what='deaths',age=seq(100),
 year=1980:2010)

R demo of explanatory plots

Plot mortality rates, population-at-risk, death counts for any age group and year

>insp.dd(dd.cmi.pens,age=50:80,year=1985:1990)
>insp.dd(dd.cmi.pens,what='pop',age=70:100,year=1988:1993)
>insp.dd(dd.cmi.pens,what='deaths',age=seq(100),
year=1980:2010)

Produce simple plots (i.e., without legend) of log- or untransformed rates:

>plot(dd.cmi.pens)
>plot(dd.cmi.pens,transform=FALSE)

1 Produce annotated plots of log or original rates:

>plot_dd(dd.cmi.pens, xlim=c(40, 110),
lpar = list(x.int = -0.2, y.int = 0.9, cex = 0.85))
>plot_dd(dd.cmi.pens, year=1985:1995, transform=FALSE)
>plot_dd(dd.cmi.pens, year=1995:1997, transform=FALSE,
lty=1:3, col=1:3)

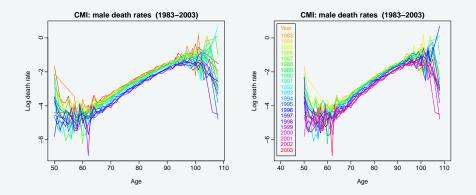
1 Produce annotated plots of log or original rates:

>plot_dd(dd.cmi.pens, xlim=c(40, 110),
lpar = list(x.int = -0.2, y.int = 0.9, cex = 0.85))
>plot_dd(dd.cmi.pens, year=1985:1995, transform=FALSE)
>plot_dd(dd.cmi.pens, year=1995:1997, transform=FALSE,
lty=1:3, col=1:3)

2 Deal with missing data

```
# without correction of empty cells
>tmp.d = extract.deaths(dd.cmi.pens, ages=55:100)
# empty cells are filled using perk model
>tmp.d = extract.deaths(dd.cmi.pens, ages=55:100,
fill='perks')
```

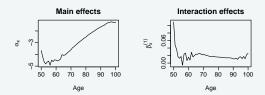
Explanatory plots: dealing with missing values



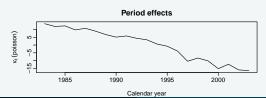
R demo for estimation of LC model

Estimate the base LC model with Poisson errors

>mod6 = lca.rh(dd.cmi.pens, mod = 'lc', interpolate=TRUE)
>coef(mod6); plot(mod6)
>fitted_plot(mod6); residual_plot(mod6)



Standard LC Regression for CMI [male]

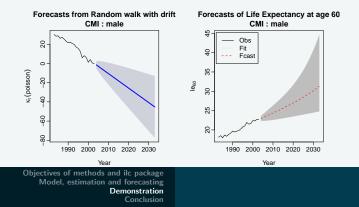


Objectives of methods and ilc package Model, estimation and forecasting Demonstration Conclusion

16 / 26

Plots of fitted models

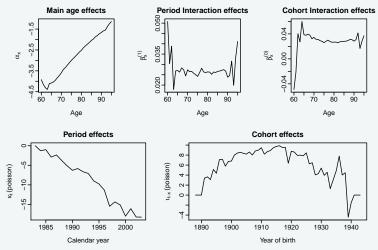
```
>forc6 = forecast(mod6, h = 20, jump = 'fit', level = 90,
shift=FALSE)
>plot_dd(forc6, xlim=c(45,100), lpar=list(x.int=-0.2,
y.int=0.9, cex=0.95))
>le6 = life.expectancy(forc6, age=60)
>flc.plot(mod6, at=60, h=30, level=90)
```



17 / 26

```
>mod1 = lca.rh(dd.cmi.pens, age=60:95, mod = "m",
restype='deviance', dec.conv=3)
>coef(mod1)
>plot(mod1)
```

Age-period-cohort plot



Age-Period-Cohort LC Regression for CMI [male]

R demo for stratified LC model

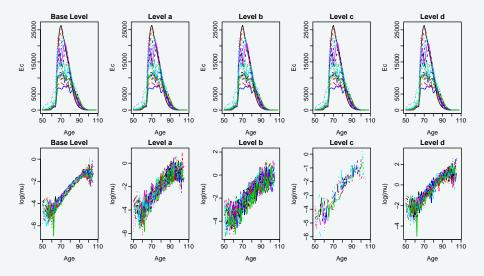
For stratified LC model, ilc package introduces a special class of data object that holds information about the grouping factors and aggregate data of number of deaths, central exposures and mortality rates

R demo for stratified LC model

- For stratified LC model, ilc package introduces a special class of data object that holds information about the grouping factors and aggregate data of number of deaths, central exposures and mortality rates
- **2** Taking the CMI experience as the base data, produce a randomly stratified mortality data

>rfp.cmi = dd.rfp(dd.cmi.pens, rfp = c(0.5,1.2,-0.7,2.5))
>matplot(rfp.cmi\$age, rfp.cmi\$pop[,,1], type='l',
xlab='Age', ylab='Ec', main = 'Base Level')
>matplot(rfp.cmi\$age, rfp.cmi\$pop[,,2], type='l',
xlab='Age', ylab='Ec', main = 'Base Level')

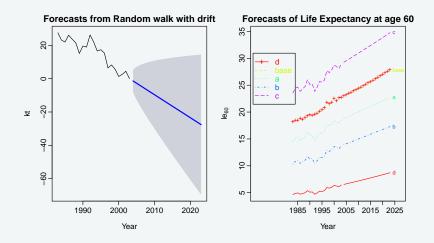
Plots of stratified Lee-Carter



R demo for estimating and forecasting of stratified LC model

>rfp.cmi = dd.rfp(dd.cmi.pens, rfp = c(0.5, 1.2, -0.7, 2.5))
>mod6e = elca.rh(rfp.cmi, age=50:100, interpolate=TRUE,
dec.conv=3, verbose=TRUE)
>coef(mod6e)
>mod6ef = forecast.lca(mod6e, h = 20, level = 90, jump='fit',
shift=FALSE)
>plot(mod6ef\$kt, ylab='kt', xlab='Year')
>matfle.plot(mod6e\$lca, mod6, at=60, label='RFP CMI', h=20)

Plot of forecast life expectancy



1 ilc package implements the Lee-Carter model with Gaussian and Poisson errors

- **1** ilc package implements the Lee-Carter model with Gaussian and Poisson errors
- 2 ilc package implements additional five models discussed in Renshaw and Haberman (2006)

- ilc package implements the Lee-Carter model with Gaussian and Poisson errors
- 2 ilc package implements additional five models discussed in Renshaw and Haberman (2006)
- **3** Stratified Lee-Carter model allows users to include additional covariates (other than age and time)

Renshaw, A. E., and S. Haberman, 2003. Lee-Carter mortality forecasting with age-specific enhancement. *Insurance: Mathematics* and *Economics*, 33, 255-272.

- Renshaw, A. E., and S. Haberman, 2003. Lee-Carter mortality forecasting with age-specific enhancement. *Insurance: Mathematics* and *Economics*, 33, 255-272.
- Hyndman, R. J., and H. L. Shang, 2010. Rainbow plots, bagplots and boxplot for functional data, *Journal of Computational and Graphical Statistics*, 19(1), 29-45.

- Renshaw, A. E., and S. Haberman, 2003. Lee-Carter mortality forecasting with age-specific enhancement. *Insurance: Mathematics* and *Economics*, 33, 255-272.
- Hyndman, R. J., and H. L. Shang, 2010. Rainbow plots, bagplots and boxplot for functional data, *Journal of Computational and Graphical Statistics*, 19(1), 29-45.
- H. L. Shang, 2011. rainbow: An R package for visualizing functional time series, *The R Journal*, 3(2), 54-59.

Thank you! · Tak

