



RLRsim: Testing for Random Effects or Nonparametric Regression Functions in Additive Mixed Models

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joint work with Sonja Greven^{1,2} and Helmut Küchenhoff¹

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Outline

Background & Problem Description

Implementation & Application Examples

Simulation Study



Linear Mixed Models

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \sum_{l=1}^L \mathbf{Z}_l \mathbf{b}_l + \boldsymbol{\varepsilon}$$

$$\mathbf{b}_l \sim \mathcal{N}_{K_l}(\mathbf{0}, \lambda_l \sigma_\varepsilon^2 \boldsymbol{\Sigma}_l), \quad \mathbf{b}_l \perp \mathbf{b}_s \quad \forall l \neq s$$

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We want to test

$$\begin{aligned} H_{0,l} : \lambda_l = 0 & \text{ versus} & H_{A,l} : \lambda_l > 0 \\ \Leftrightarrow H_{0,l} : \text{Var}(\mathbf{b}_l) = 0 & \text{ versus} & H_{A,l} : \text{Var}(\mathbf{b}_l) > 0 \end{aligned}$$

Application examples:

- ▶ testing for equality of means between groups/subjects
- ▶ testing for linearity of a smooth function



Additive Models as Linear Mixed Models

Simple additive model:

$$\mathbf{y} = f(\mathbf{x}) + \varepsilon$$
$$f(x_i) \approx \sum_{j=1}^J \delta_j B_j(x_i)$$

- ▶ fit via PLS: $\min_{\delta} (\|\mathbf{y} - \mathbf{B}\delta\|^2 + \frac{1}{\lambda} \delta' \mathbf{P} \delta)$
- ▶ reparametrize s.t. PLS-estimation is equivalent to (RE)ML-estimation



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- ▶ reparametrize s.t. PLS-estimation is equivalent to (RE)ML-estimation given λ in a LMM with
 - ▶ fixed effects for the *unpenalized* part of $f(\mathbf{x})$
 - ▶ random effects ($\overset{\text{i.i.d.}}{\sim} \mathcal{N}(0, \lambda \sigma_{\varepsilon}^2)$) for the *deviations from the unpenalized* part

(Brumback, Ruppert, Wand, 1999; Fahrmeir, Kneib, Lang, 2004)

- ▶ In R: `mgcv::gamm()`, `lmeSplines`



Problem:

Likelihood Ratio Tests for Zero Variance Components

General Case:

- ▶ $y_1, \dots, y_n \stackrel{\text{i.i.d.}}{\sim} f(y|\boldsymbol{\theta}); \boldsymbol{\theta} = (\theta_1, \dots, \theta_p)$
- ▶ Test: $H_0 : \theta_i = \theta_i^0$ versus $H_A : \theta_i \neq \theta_i^0$
- ▶ $LRT = 2 \log L(\hat{\boldsymbol{\theta}}|\mathbf{y}) - 2 \log L(\hat{\boldsymbol{\theta}}^0|\mathbf{y}) \stackrel{n \rightarrow \infty}{\sim} \chi_1^2$



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Problem for testing $H_0 : \text{Var}(\mathbf{b}_1) = 0$

Underlying assumptions for asymptotics violated:

- ▶ data in LMM not independent
- ▶ $\boldsymbol{\theta}^0$ not an interior point of the parameter space Θ



Previous Results:

- ▶ **Stram, Lee (1994); Self, Liang (1987):** for i. i. d. observations/subvectors, testing on the boundary of Θ :
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- ▶ **Crainiceanu, Ruppert (2004):**
 - ▶ Stram/Lee mixture very conservative for non-i. i. d. data, small samples
 - ▶ LRT often with large point mass at zero, restricted LRT ($RLRT$) more useful
 - ▶ derive exact finite sample distributions of LRT and $RLRT$ in LMMs with *one* variance component



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- ▶ **Greven et al. (2007):**
pseudo-ML arguments to justify application of results in Crainiceanu, Ruppert (2004) to models with multiple variance components



RLRsim: Algorithm

$$RLRT_n \sim \sup_{\lambda \geq 0} \left((n-p) \log \left(1 + \frac{N_n(\lambda)}{D_n(\lambda)} \right) - \sum_{k=1}^K \log (1 + \lambda \mu_{k,n}) \right),$$

$$N_n(\lambda) = \sum_{k=1}^K \frac{\lambda \mu_{k,n}}{1 + \lambda \mu_{k,n}} w_k^2; \quad D_n(\lambda) = \sum_{k=1}^K \frac{w_k^2}{1 + \lambda \mu_{k,n}} + \sum_{k=K+1}^{n-p} w_k^2$$

$w_k \sim \mathcal{N}(0, 1)$; μ : eigenvalues of $\Sigma^{1/2} \mathbf{Z}' (\mathbf{I}_n - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}) \mathbf{Z} \Sigma^{1/2}$



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Rapid simulation from this distribution:

- ▶ do eigenvalue decomposition to get μ
- ▶ repeat:
 - ▶ draw $(K+1)$ χ^2 variates
 - ▶ one-dimensional maximization in λ (via grid search)

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Rapid simulation from this distribution:

- ▶ do eigenvalue decomposition to get μ
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→ computational cost depends on K , not n

→ implemented in C \Rightarrow quasi-instantaneous

→ easy extension to models with $L > 1$



Example: One Variance Component

Test for random intercept (`nlme::lme`):

```
> m0 <- lme(distance ~ age + Sex, data = Orthodont, random = ~ 1)
> system.time(print( exactRLRT(m0) ), gcFirst=T)
```

```
simulated finite sample distribution of RLRT.
(p-value based on 10000 simulated values)
```

```
RLRT = 47.0114, p-value < 2.2e-16
```

user	system	elapsed
0.42	0.00	0.42

```
> system.time(simulate.lme(m0,nsim=10000,method='REML'), gcFirst=T)
```

user	system	elapsed
55.00	0.03	55.48



Example: Two Variance Components

Test for random slope with nuisance random intercept

(`lme4::lmer`):

```
> m0 <- lmer(distance ~ age + Sex + (1|Subject), data = Orthodont)
> mA <- update(m0, .~. + (0 + age|Subject))
> mSlope <- update(mA, .~. - (1|Subject))
> exactRLRT(mSlope, mA, m0)
```

simulated finite sample distribution of RLRT.

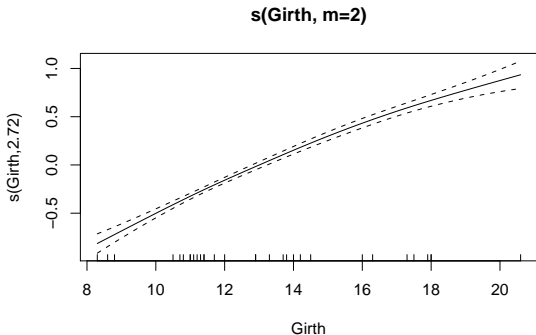
(p-value based on 10000 simulated values)

RLRT = 0.8672, p-value = 0.1603



Example: Testing for Linearity of a Smooth Function

```
> library(mgcv); data(trees)
> m1 <- gamm(I(log(Volume)) ~ Height + s(Girth, m = 2),
+           data = trees)$lme
```



Significant deviations from linearity?



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> library(mgcv); data(trees)
> m1 <- gamm(I(log(Volume)) ~ Height + s(Girth, m = 2),
+           data = trees)$lme
> exactRLRT(m1)
```

simulated finite sample distribution of RLRT.

(p-value based on 10000 simulated values)

RLRT = 5.4561, p-value = 0.0052

Simulation Study: Settings

H_0	tested VC	nuisance VCs
equality of group means	random intercept	- random slope uni-/bivariate smooth
equality of group trends	random slope	random intercept
no effect / linearity	univariate smooth	- random intercept uni-/bivariate smooth
additivity	bivariate smooth	2 univariate smooths

Goal: compare size & power of tests for zero variance components

- ▶ sample sizes $n = 50, 100, 500$
- ▶ mildly unbalanced group sizes for $K = 5, 20$
- ▶ details: Scheipl, Greven, Küchenhoff (2007)



Simulation study

Compared Tests:

- ▶ *RLR*-type tests:
RLRsim, parametric bootstrap, $0.5\delta_0 : 0.5\chi_1^2$
- ▶ *F*-type tests:
bootstrap *F*-type statistics, mgcv's approximate *F*-test,
SAS-implementations of generalized *F*-test etc..



Simulation study

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Main Results:

- ▶ RLRsim: equivalent performance to bootstrap RLRT, but practically instantaneous
- ▶ χ^2 -mixture approximation for RLRT: always conservative, lower than nominal size & reduced power
- ▶ bootstrap RLRT, bootstrap *F*-type statistics similar
- ▶ *F*-test from mgcv: similar power as χ^2 -mixture, occasionally seriously anti-conservative



Conclusion

- ▶ conventional RLRTs for $\text{Var}(\text{Random Effect}) = 0$ are broken, but not beyond repair.

⇒ `RLRsim`

- ▶ is a rapid, more powerful alternative that performs as well as a parametric bootstrap.
- ▶ has a convenient interface for models fit with `nlme::lme` or `lme4::lmer`.
- ▶ Current limitations: no correlated random effects, no serial correlation, only Gaussian responses.



Further Reading:

- ▶ Crainiceanu, C. and Ruppert, D. (2004) *Likelihood ratio tests in linear mixed models with one variance component*, JRSS-B, **66**(1), 165–185.
- ▶ Greven, S., Crainiceanu, C. M., Küchenhoff, H. and Peters, A. (2008) *Restricted Likelihood Ratio Testing for Zero Variance Components in Linear Mixed Models*, JCGS, to appear.
- ▶ Scheipl, F., Greven, S., and Küchenhoff, H. (2008) *Size and power of tests for a zero random effect variance or polynomial regression in additive and linear mixed models*, CSDA, **52**(7), 3283–3299.