Regularization Paths

Trevor Hastie

Stanford University

drawing on collaborations with Brad Efron, Mee-Young Park, Saharon

Rosset, Rob Tibshirani, Hui Zou and Ji Zhu.



- Boosting fits a regularization path toward a max-margin classifier. Sympath does as well.
- In neither case is this endpoint always of interest somewhere along the path is often better.
- Having efficient algorithms for computing entire paths facilitates this selection.
- A mini industry has emerged for generating regularization paths covering a broad spectrum of statistical problems.

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Boosting Stumps for Regression



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Linear Regression

Here is a version of least squares boosting for multiple linear regression: (assume predictors are standardized)

(Incremental) Forward Stagewise

- 1. Start with $r = y, \beta_1, \beta_2, \dots, \beta_p = 0$.
- 2. Find the predictor x_i most correlated with r
- 3. Update $\beta_i \leftarrow \beta_i + \delta_i$, where $\delta_i = \epsilon \cdot \operatorname{sign} \langle r, x_i \rangle$
- 4. Set $r \leftarrow r \delta_j \cdot x_j$ and repeat steps 2 and 3 many times

 $\delta_i = \langle r, x_i \rangle$ gives usual forward stagewise; different from forward stepwise

Analogous to least squares boosting, with *trees=predictors*



Coefficients

0.0

0.5

 $\mathbf{t} = \sum_{j}^{1.0} |\beta_j|$

2.5

2.0

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Example: Prostate Cancer Data

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Forward Stagewise Lasso Icavol 0.6 0.6 0.4 0.4 Coefficients reiah pgg45 0.2 0.2 0.0 0.0 age age -0.2 -0.2

100 150 200

Iteration

0 50

Linear regression via the Lasso (Tibshirani, 1995)

- Assume $\bar{y} = 0$, $\bar{x}_i = 0$, $\operatorname{Var}(x_i) = 1$ for all j.
- Minimize $\sum_i (y_i \sum_j x_{ij}\beta_j)^2$ subject to $||\beta||_1 \leq t$
- Similar to *ridge regression*, which has constraint $||\beta||_2 \leq t$
- Lasso does variable selection and shrinkage, while ridge only shrinks.



Least Squares Boosting

Friedman, Hastie & Tibshirani — see *Elements of Statistical* Learning (chapter 10)

Supervised learning: Response y, predictors $x = (x_1, x_2 \dots x_p)$.

- 1. Start with function F(x) = 0 and residual r = y
- 2. Fit a CART regression tree to r giving f(x)
- 3. Set $F(x) \leftarrow F(x) + \epsilon f(x), r \leftarrow r \epsilon f(x)$ and repeat steps 2 and 3 many times

Diabetes Data

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$\label{eq:like} Like\ a\ ``more\ democratic''\ version\ of\ forward\ stepwise\ regression.$

- 1. Start with $r = y, \hat{\beta}_1, \hat{\beta}_2, \dots \hat{\beta}_p = 0$. Assume x_j standardized.
- 2. Find predictor x_i most correlated with r.
- 3. Increase β_j in the direction of sign(corr (r, x_j)) until some other competitor x_k has as much correlation with current residual as does x_j .
- 4. Move $(\hat{\beta}_j, \hat{\beta}_k)$ in the joint least squares direction for (x_j, x_k) until some other competitor x_ℓ has as much correlation with the current residual
- 5. Continue in this way until all predictors have been entered. Stop when $\operatorname{corr}(r, x_j) = 0 \forall j$, i.e. OLS solution.



df for LAR

- *df* are labeled at the top of the figure
- At the point a competitor enters the active set, the *df* are incremented by 1.
- Not true, for example, for stepwise regression.

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The LAR direction \mathbf{u}_2 at step 2 makes an equal angle with \mathbf{x}_1 and \mathbf{x}_2 .

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- Relationship between the 3 algorithms
- Lasso and forward stagewise can be thought of as restricted versions of LAR
- *Lasso*: Start with LAR. If a coefficient crosses zero, stop. Drop that predictor, recompute the best direction and continue. This gives the Lasso path

Proof: use KKT conditions for appropriate Lagrangian. Informally:

$$\frac{\partial}{\partial \beta_j} \left[\frac{1}{2} ||\mathbf{y} - \mathbf{X}\beta||^2 + \lambda \sum_j |\beta_j| \right] = 0$$

$$\Leftrightarrow$$

$$\langle \mathbf{x}_j, \mathbf{r} \rangle = \lambda \cdot \operatorname{sign}(\hat{\beta}_j) \quad \text{if } \hat{\beta}_j \neq 0 \text{ (active)}$$

- Forward Stagewise: Compute the LAR direction, but constrain the sign of the coefficients to match the correlations $corr(r, x_j)$.
- The incremental forward stagewise procedure approximates these steps, one predictor at a time. As step size $\epsilon \to 0$, can show that it coincides with this modified version of LAR

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Cross-Validation Error Curve

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lars package

- The LARS algorithm computes the entire Lasso/FS/LAR path in same order of computation as one full least squares fit.
- When $p \gg N$, the solution has at most N non-zero coefficients. Works efficiently for micro-array data (p in thousands).
- Cross-validation is quick and easy.



- 10-fold CV error curve using lasso on some diabetes data (64 inputs, 442 samples).
- Thick curve is CV error curve
- Shaded region indicates standard error of CV estimate.
- Curve shows effect of overfitting — errors start to increase above s = 0.2.
- This shows a trade-off between bias and variance.

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Forward Stagewise and the Monotone Lasso





- - Expand the variable set to include their negative versions $-x_i$.
 - Original lasso corresponds to a *positive* lasso in this enlarged space.
 - Forward stagewise corresponds to a *monotone lasso*. The L_1 norm $||\beta||_1$ in this enlarged space is *arc-length*.
 - Forward stagewise produces the maximum decrease in loss per unit arc-length in coefficients.

Degrees of Freedom of Lasso

- The df or effective number of parameters give us an indication of how much fitting we have done.
- Stein's Lemma: If y_i are i.i.d. $N(\mu_i, \sigma^2)$,

$$df(\hat{\boldsymbol{\mu}}) \stackrel{\text{def}}{=} \sum_{i=1}^{n} \operatorname{cov}(\hat{\mu}_i, y_i) / \sigma^2 = E\left[\sum_{i=1}^{n} \frac{\partial \hat{\mu}_i}{\partial y_i}\right]$$

- Degrees of freedom formula for LAR: After k steps, $df(\hat{\mu}_k) = k$ exactly (amazing! with some regularity conditions)
- Degrees of freedom formula for lasso: Let $\hat{df}(\hat{\mu}_{\lambda})$ be the number of *non-zero* elements in $\hat{\beta}_{\lambda}$. Then $E\hat{d}f(\hat{\mu}_{\lambda}) = df(\hat{\mu}_{\lambda})$.

0

200

500

0.0

ad Coeffic



4 5

2 3

0.2

0.4

|beta|/max|beta|

0.6

0.8

df for LAR

- *df* are labeled at the top of the figure
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- Not true, for example, for stepwise regression.

Back to Boosting

- Work with Rosset and Zhu (JMLR 2004) extends the connections between Forward Stagewise and L₁ penalized fitting to other loss functions. In particular the Exponential loss of Adaboost, and the Binomial loss of Logitboost.
- In the separable case, L_1 regularized fitting with these losses converges to a L_1 maximizing margin (defined by β^*), as the penalty disappears. i.e. if

$$\beta(t) = \arg \min L(y, f) \quad \text{s.t.} \ |\beta| \le t,$$

then

$$\lim_{t\uparrow\infty}\frac{\beta(t)}{|\beta(t)|}\to\beta^*$$

• Then $\min_i y_i F * (x_i) = \min_i y_i x_i^T \beta^*$, the L_1 margin, is maximized.

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- When the monotone lasso is used in the expanded feature space, the connection with boosting (with shrinkage) is more precise.
- This ties in very nicely with the L_1 margin explanation of boosting (Schapire, Freund, Bartlett and Lee, 1998).
- makes connections between SVMs and Boosting, and makes explicit the margin maximizing properties of boosting.
- experience from statistics suggests that some $\beta(t)$ along the path might perform better—a.k.a stopping early.
- Zhao and Yu (2004) incorporate backward corrections with forward stagewise, and produce a boosting algorithm that mimics lasso.

Maximum Margin and Overfitting

Mixture data from ESL. Boosting with 4-node trees, gbm package in R, shrinkage = 0.02, Adaboost loss.



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Lasso or Forward Stagewise?

- Micro-array example (Golub Data). N = 38, p = 7129,response binary ALL vs AML
- Lasso behaves chaotically near the end of the path, while Forward Stagewise is smooth and stable.



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- Bach and Jordan (2004) have path algorithms for Kernel estimation, and for efficient ROC curve estimation. The latter is a useful generalization of the Sympath algorithm discussed later.
- Rosset and Zhu (2004) discuss conditions needed to obtain piecewise-linear paths. A combination of piecewise quadratic/linear loss function, and an L_1 penalty, is sufficient.
- Mee-Young Park is finishing a *Cosso* path algorithm. Cosso (Lin and Zhang, 2002) fits models of the form

$$\min_{\beta} \ell(\beta) + \sum_{k=1}^{K} \lambda_k ||\beta_k||_2$$

where $|| \cdot ||_2$ is the L_2 norm (not squared), and β_k represents a *subset* of the coefficients.

Other Path Algorithms

- *Elasticnet:* (Zou and Hastie, 2005). Compromise between lasso and ridge: minimize $\sum_i (y_i - \sum_j x_{ij}\beta_j)^2$ subject to $\alpha ||\beta||_1 + (1 - \alpha) ||\beta||_2^2 \le t$. Useful for situations where variables operate in correlated groups (genes in pathways).
- *Glmpath:* (Park and Hastie, 2005). Approximates the L_1 regularization path for *generalized linear models*. e.g. logistic regression, Poisson regression.
- Friedman and Popescu (2004) created *Pathseeker*. It uses an efficient incremental forward-stagewise algorithm with a variety of loss functions. A generalization adjusts the leading k coefficients at each step; k = 1 corresponds to forward stagewise, k = p to gradient descent.

elasticnet package (Hui Zou)

- $\operatorname{Min}\sum_{i}(y_i \sum_{j} x_{ij}\beta_j)^2$ s.t. $\alpha \cdot ||\beta||_2^2 + (1 \alpha) \cdot ||\beta||_1 \le t$
- Mixed penalty selects correlated sets of variables in *groups*.
- For fixed α, LARS algorithm, along with a standard *ridge regression* trick, lets us compute the entire regularization path.





glmpath package

- $\max \ell(\beta)$ s.t. $||\beta||_1 \le t$
- Predictor-corrector methods in convex optimization used.
- Computes exact path at a sequence of index
- Can approximate the junctions (in t) where the active set changes.
- coxpath included in package.

Path algorithms for the SVM

- The two-class SVM classifier $f(X) = \alpha_0 + \sum_{i=1}^N \alpha_i K(X, x_i) y_i$ can be seen to have a quadratic penalty and piecewise-linear loss. As the cost parameter C is varied, the *Lagrange multipliers* α_i change piecewise-linearly.
- This allows the entire regularization path to be traced exactly. The active set is determined by the points exactly on the margin.



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This *hinge loss* criterion is equivalent to the SVM, with λ monotone in B. Compare with

$$\min_{\beta_0, \beta} \sum_{i=1}^N \log \left[1 + e^{-y_i f(x_i)} \right] + \frac{\lambda}{2} \|\beta\|^2$$

This is *binomial deviance loss*, and the solution is "ridged" linear logistic regression.



- γ is a kernel parameter: $K(x, z) = \exp(-\gamma ||x z||^2)$.
- λ (or C) are regularization parameters, which have to be determined using some means like cross-validation.

JMLR 2004)

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Concluding Comments

- Using logistic regression + binomial loss or Adaboost exponential loss, and same quadratic penalty as SVM, we get the same limiting margin as SVM (Rosset, Zhu and Hastie,
- Alternatively, using the "Hinge loss" of SVMs and an L₁ penalty (rather than quadratic), we get a *Lasso* version of SVMs (with at most N variables in the solution for any value of the penalty.

- Boosting fits a monotone L₁ regularization path toward a maximum-margin classifier
- Many modern function estimation techniques create a path of solutions via regularization.
- In many cases these paths can be computed efficiently and entirely.
- This facilitates the important step of model selection selecting a desirable position along the path using a test sample or by CV.